# EE 505 Lecture 10

Statistical Circuit Modeling

Review from previous lecture:

## Summary of Results

Structure	Nominal Resistance	Standard Deviation	Normalized Standard Deviation
R	R <sub>N</sub>	$\sigma_{R_{R}}$	$\sigma_{rac{R_R}{R_N}}$
Ser nR	nR <sub>N</sub>	$\sqrt{n}\sigma_{R_R}$	$rac{1}{\sqrt{n}}\sigma_{rac{R_R}{R_N}}$
Par nR	$\frac{R_{N}}{n}$	$rac{1}{n^{3/2}}\sigma_{\scriptscriptstyle R_{\scriptscriptstyle R}}$	$rac{1}{\sqrt{n}}\sigma_{rac{R_R}{R_N}}$
Ser 2R Par 2R	$2R_N$ $\frac{R_N}{2}$	$\sqrt{2}\sigma_{\rm R_R}$ $\sigma_{\rm R_R}$ $\sqrt{8}$	$\sigma_{rac{R_R}{R_N}}/\sqrt{2}$ $\sigma_{rac{R_R}{R_N}}/\sqrt{2}$
Ser 4R	4R <sub>N</sub>	$2\sigma_{_{R_{P}}}$	$\sigma_{\frac{R_{R}}{R_{N}}}$ 2
Par 4R	$\frac{R_{N}}{4}$	$2\sigma_{R_{R}}$	$\sigma_{\frac{R_{R}}{R_{N}}}$ 2 $\sigma_{\frac{R_{R}}{R_{N}}}$ 2
Par/Ser 4R	R <sub>N</sub>	$\sigma_{_{R_{R}}}$ 2	$\frac{r_{R}}{R_{N}}$ 2

Review from previous lecture:

Consider a resistor of width W and length L  $\sigma_{R}^{2} = \left(\frac{L}{W}\right)^{2} \bullet \frac{\sigma_{REF}^{2}}{W \bullet L} = \sigma_{REF}^{2} \bullet \frac{L}{W^{3}} \qquad A=W\bullet L$ Note  $\sigma_{R}$  is dependent on resistance value Consider now the normalized resistance  $\frac{R}{R_{N}}$ where  $R_{N} = R_{GN} \frac{L}{W}$ 

It follows that

$$\sigma_{\frac{R}{R_{N}}}^{2} = \left(\frac{1}{R_{N}^{2}}\right) \left(\sigma_{REF}^{2} \frac{L}{W^{3}}\right) = \left(\frac{W^{2}}{R_{\square N}^{2}L^{2}}\right) \left(\sigma_{REF}^{2} \frac{L}{W^{3}}\right) = \left(\frac{1}{WL}\right) \left[\frac{\sigma_{REF}^{2}}{R_{\square N}^{2}}\right]$$

The term on the right in [] is the ratio of two process parameters so define the process parameter  $A_R$  by the expression  $A_R = \frac{\sigma_{REF}}{R_{_{IN}}}$ 

 $A_R$  is more convenient to use than both  $\sigma_{REF}$  and  $R_{_{\square}N}$ 

Thus the normalized resistance is given by the expression

$$\sigma_{\frac{R}{R_{N}}}^{2} = \frac{A_{R}^{2}}{WL} = \frac{A_{R}^{2}}{A}$$

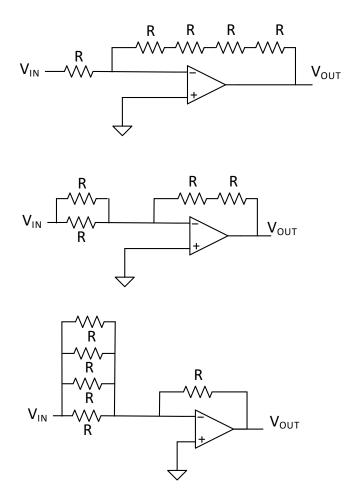
Note  $\sigma_{R/RN}$  is not dependent on resistance value Will term  $A_R$  the "Pelgrom parameter" (though Pelgrom only presented results for MOS devices)

### **Amplifier Gain Accuracy**

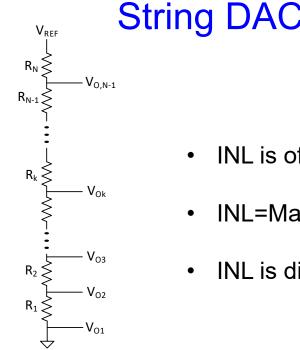
Review from previous lecture:

### **Amplifier Gain Accuracy**

Many different ways to achieve a given gain with a given resistor area



Which will have the best yield?



 $0 \le k \le N-1$ 

- INL is of considerable interest
- INL=Max(|INL<sub>k</sub>|), 0<k<N-1
- INL is difficult to characterize analytically so will focus on INL<sub>k</sub>

Assume resistors are uncorrelated RVs but identically distributed, typically zero mean Gaussian

Consider INL<sub>k</sub>= 
$$V_{OUT}(k) - V_{FIT}(k)$$

$$V_{OUT}(k) = \begin{cases} 0 & k = 0\\ \sum_{j=1}^{k} R_{j} & 1 \le k \le N-1\\ \sum_{j=1}^{N} R_{j} & 1 \le k \le N-1 \end{cases}$$
$$V_{FIT}(k) = \frac{k}{N-1} \frac{\sum_{j=1}^{N-1} R_{j}}{\sum_{j=1}^{N} R_{j}} V_{REF} & 0 \le k \le N-1 \end{cases}$$

$$INL_{k} = \frac{\left(\sum_{j=1}^{k} R_{j} - \frac{k}{N-1} \sum_{j=1}^{N-1} R_{j}\right)}{\frac{V_{REF}}{2^{n}}} \qquad 1 \le k \le N-1$$

$$INL_{k} = \frac{\sum_{j=1}^{k} R_{j} - \frac{k}{N-1} \sum_{j=1}^{N-1} R_{j}}{\sum_{j=1}^{N} R_{j}} 2^{n} \qquad 1 \le k \le N-1$$

$$INL_{k} = \frac{\sum_{j=1}^{k} R_{j} - \frac{k}{N-1} \sum_{j=1}^{k} R_{j} - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{j}}{\sum_{j=1}^{N} R_{j}} 2^{n} \qquad 1 \le k \le N-1$$

$$INL_{k} = \frac{\sum_{j=1}^{k} R_{j} \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{j}}{\sum_{j=1}^{N} R_{j}} 2^{n} \qquad 1 \le k \le N-1$$

Let  $R_j = R_{NOM} + R_{Rj}$ 

$$INL_{k} = \frac{\left[\sum_{j=1}^{k} R_{NOM} \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{NOM}\right] + \sum_{j=1}^{k} R_{Rj} \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj}}{\sum_{j=1}^{N} R_{NOM} + \sum_{j=1}^{N} R_{Rj}} 2^{n} \qquad 1 \le k \le N-1$$

$$INL_{k} = \frac{R_{NOM} \left[k \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \left(N-k-1\right)\right] + \sum_{j=1}^{k} R_{Rj} \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj}}{NR_{NOM} + \sum_{j=1}^{N} R_{Rj}} 2^{n} \qquad 1 \le k \le N-1$$

$$INL_{k} = \frac{2^{n}}{NR_{NOM}} \frac{\sum_{j=1}^{k} R_{Rj} \left(1 - \frac{k}{N-1}\right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj}}{1 + \frac{1}{NR_{NOM}} \sum_{j=1}^{N} R_{Rj}} \qquad 1 \le k \le N-1$$

1

If we do a Taylor's series expansion of the reciprocal of the denominator and eliminate second-order and higher terms it follows that

$$INL_{k} = \frac{1}{R_{NOM}} \left[ \sum_{j=1}^{k} R_{Rj} \left( 1 - \frac{k}{N-1} \right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj} \right] \left[ 1 - \frac{1}{NR_{NOM}} \sum_{j=1}^{N} R_{Rj} \right] \qquad 1 \le k \le N - 1$$
$$INL_{k} = \frac{1}{R_{NOM}} \left[ \sum_{j=1}^{k} R_{Rj} \left( 1 - \frac{k}{N-1} \right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj} \right] \qquad 1 \le k \le N - 1$$

Note that INK<sub>k</sub> is a zero-mean multivariate Gaussian distribution

$$INL_{k} = \frac{1}{R_{NOM}} \left[ \sum_{j=1}^{k} R_{Rj} \left( 1 - \frac{k}{N-1} \right) - \frac{k}{N-1} \sum_{j=k+1}^{N-1} R_{Rj} \right] \qquad 1 \le k \le N-1$$

Since the resistors are identically distributed and the coefficients are not a function of the index j, it follows that

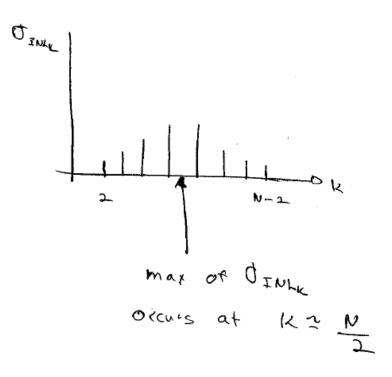
$$\sigma_{INLk}^{2} = \sigma_{\frac{R_{R}}{R_{NOM}}}^{2} \left[ \sum_{j=1}^{k} \left( 1 - \frac{k}{N-1} \right)^{2} + \sum_{j=k+1}^{N-1} \left( \frac{k}{N-1} \right)^{2} \right] \qquad 1 \le k \le N-1$$

Since the index in the sum does not appear in the arguments, this simplifies to

$$\sigma_{INLk} = \sigma_{\frac{R_R}{R_{NOM}}} \sqrt{\frac{(N-1-k)k}{N-1}} \qquad 1 \le k \le N-1$$

Note there is a nice closed-form expression for the INL<sub>k</sub> for a string DAC !!

INL<sub>k</sub> assumes a maximum variance at mid-code



$$\sigma_{INLk\max} = \sigma_{\frac{R_R}{R_{NOM}}} \frac{\sqrt{N}}{2}$$

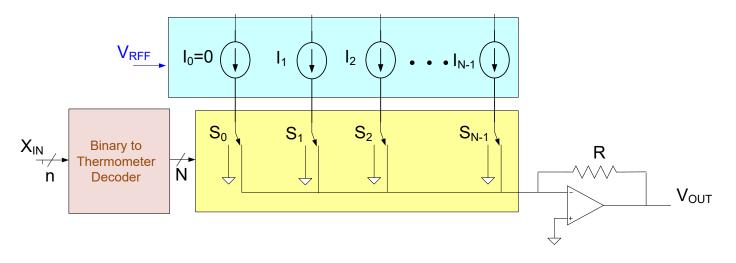
How about statistics for the INL?

$$INL_{k} = \sum_{i=1}^{k-1} \frac{R_{R_{i}}}{R_{N}} \left( \frac{N-k}{N-i} \right) - \sum_{i=k}^{N-1} \frac{R_{R_{i}}}{R_{N}} \left( \frac{k-i}{N-i} \right)$$

INL is an order statistic

Distribution functions for order statistics are very complicated and closed form solutions do not exist

INL is not zero-mean and not Gaussian



Assume unary current source array and define  $I_0=0$ 

$$V_{OUT}(k) = -R \sum_{j=0}^{k-1} I_j \qquad 1 \le k \le N$$

For notational convenience will normalize by -R to obtain

$$I_{OUTX}(k) = \sum_{i=0}^{n-1} I_i \qquad 1 \le k \le N$$

Assume current sources are random variables with identical distributions

$$I_{j} = I_{NOM} + I_{Rj} \qquad \qquad I_{Rj} \propto N(0, \sigma_{I})$$

$$INL_{k}(k) = \frac{\sum_{j=0}^{k-1} I_{j} - I_{FIT}(k)}{I_{NOM}} \qquad 1 \le k \le N$$

$$X_{N} \xrightarrow{+} n$$
Binary to  
Thermometer  
Binary to  
Thermometer  
Binary to  
Thermometer  
N
Binary to  
Thermometer  
N
Binary to  

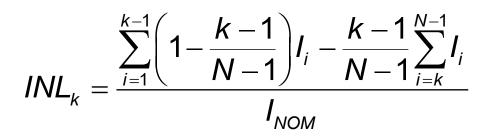
$$I_{FIT}(k) = \frac{k-1}{N-1} \left( \sum_{j=1}^{N-1} I_j \right) \qquad 1 \le$$

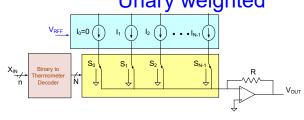
 $1 \le k \le N$ 

$$INL_{k}(k) = \frac{\sum_{j=1}^{k-1} I_{j} - \frac{k-1}{N-1} \left( \sum_{j=1}^{N-1} I_{j} \right)}{I_{NOM}}$$

$$INL_{k} = \frac{\sum_{i=1}^{k-1} \left(1 - \frac{k-1}{N-1}\right) I_{i} - \frac{k-1}{N-1} \sum_{i=k}^{N-1} I_{i}}{I_{NOM}}$$

1 T





Model the current sources as 
$$I_{j} = I_{NOM} + I_{Rj}$$
  
 $INL_{k} = \frac{\sum_{i=1}^{k-1} \left(1 - \frac{k-1}{N-1}\right) (I_{NOM} + I_{Rk}) - \frac{k-1}{N-1} \sum_{i=k}^{N-1} (I_{NOM} + I_{Rk})}{I_{NOM}}$ 

It can be shown that the nominal part cancels, thus

$$INL_{k} = \sum_{i=1}^{k-1} \left(\frac{N-k}{N-1}\right) \left(\frac{I_{Rk}}{I_{NOM}}\right) - \frac{k-1}{N-1} \sum_{i=k}^{N-1} \left(\frac{I_{Rk}}{I_{NOM}}\right)$$

This is a sum of uncorrelated random variables

The variance of  $I_{NKk}$  can be readily calculated

$$X_{IN} \xrightarrow[n]{\text{Binary to}} I_0 = 0 \xrightarrow[n]{\text{I}_1} I_2 \xrightarrow[n]{\text{I}_2} \cdots \xrightarrow[n]{\text{I}_{N-1}} I_1 \xrightarrow[n]{\text{I}_2} \xrightarrow[n]{\text$$

$$\sigma_{INL_{k}}^{2} = \sum_{i=1}^{k-1} \left( \frac{N-k}{N-1} \right)^{2} \sigma_{\frac{I_{Rk}}{I_{NOM}}}^{2} + \left( \frac{k-1}{N-1} \right)^{2} \sum_{i=k}^{N-1} \sigma_{\frac{I_{Rk}}{I_{NOM}}}^{2}$$

$$I_j = I_N + I_{Rj}$$

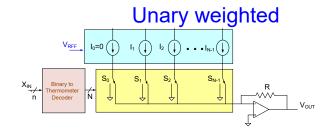
$$\sigma_{INL_{k}}^{2} = \left[ \left(k-1\right) \left(\frac{N-k}{N-1}\right)^{2} + \left(N-k\right) \left(\frac{k-1}{N-1}\right)^{2} \right] \sigma_{\frac{I_{Rk}}{I_{NOM}}}^{2}$$

This simplifies to

$$\sigma_{INL_{k}}^{2} = \frac{(k-1)(N-k)}{(N-1)}\sigma_{\frac{I_{Rk}}{I_{NOM}}}^{2}$$

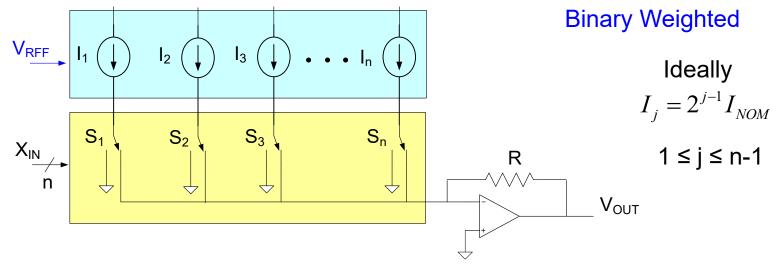
As for the string DAC, the maximum  $INL_k$  occurs near mid-code at about k=N/2 thus

$$\sigma_{INL_{k-MAX}} = \sigma_{\frac{I_R}{I_{NOM}}} \left[ \frac{\sqrt{N}}{2} \right]$$



 $I_j = I_N + I_{Rj}$ 

And, as for the string DAC, the INL is an order statistic and thus a closed-form solution does not exist



The structure looks about the same as for the unary structure but now the current sources are binary weighted

$$V_{OUT}(\mathbf{b}) = -R \sum_{j=0}^{n} b_{i} I_{j} \qquad \mathbf{b} = \langle b_{n}, b_{n-1} \dots b_{1} \rangle$$

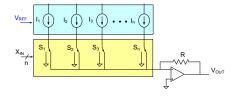
Define the decimal equivalent of b,  $k_b$ , by

$$k_{\mathbf{b}} = \sum_{j=1}^{n} b_j 2^{j-1}$$

For notational convenience will normalize by -R to obtain

$$I_{OUTX}(\mathbf{b}) = \sum_{i=1}^{n} b_i I_i$$
 for <0,0,...0>  $\leq \mathbf{b} \leq <1,1,...1>$ 

$$I_{FIT}(\mathbf{b}) = \frac{k_{\mathbf{b}}}{N-1} \sum_{i=1}^{n} I_i \qquad 0 \le k_{\mathbf{b}} \le N-1$$



Thus

$$INL_{k}(b) = rac{I_{OUTX}(b) - I_{FIT}(b)}{I_{LSBX}}$$

for  $<0,0,...0> \le b \le <1,1,...1>$ or equivalently for  $0 \le k_b \le N-1$ 

$$INL_{k}(b) = \frac{\sum_{i=1}^{n} b_{i}I_{i} - \frac{k_{b}}{N-1}\sum_{i=1}^{n}I_{i}}{I_{LSBX}}$$

Assume bundled current sources are comprised of unary current sources from same distribution

$$I_m = \sum_{k=2^{m-1}}^{2^m - 1} I_{Gk}$$
  $I_{Gk} = I_{NOM} + I_{RGk}$ 

Binary Weighted  

$$V_{RFF}$$
  $h \downarrow l_2 \downarrow l_3 \downarrow \cdots l_n \downarrow$   
 $x_{N}$   $h \downarrow s_2 \downarrow s_3 \downarrow \cdots s_n$   $x_{N}$   $y_{OUT}$ 

$$INL_{b} = \frac{\sum_{i=1}^{n} \left( b_{i} \left( \sum_{k=2^{i-1}}^{2^{i}-1} I_{Gk} \right) \right) - \frac{k_{b}}{N-1} \sum_{i=1}^{2^{n}-1} I_{Gi}}{I_{LSBX}}$$

Substituting the values for  $I_{Gk}$ , it can be shown that the nominal parts cancel thus

$$INL_{b} = \frac{\sum_{i=1}^{n} \left( b_{i} \left( \sum_{k=2^{i-1}}^{2^{i}-1} I_{RGk} \right) \right) - \frac{k_{b}}{N-1} \sum_{i=1}^{2^{n}-1} I_{RGi}}{I_{LSBX}}$$

This can be expressed as

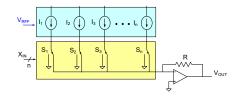
$$INL_{b} = \sum_{i=1}^{n} \sum_{k=2^{i-1}}^{2^{i}-1} \left[ b_{i} - \frac{k_{b}}{N-1} \right] \frac{I_{RGk}}{I_{LSBX}}$$

This is now a sum of uncorrelated random variables, thus

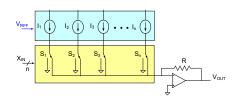
$$\sigma_{INL_{b}} = \sqrt{\sum_{i=1}^{n} \sum_{k=2^{i-1}}^{2^{i}-1} \left[ b_{i} - \frac{k_{b}}{N-1} \right]^{2}} \bullet \sigma_{\frac{I_{RGK}}{I_{LSBX}}}$$

This reduces to

$$\sigma_{INL_{b}} = \sqrt{\sum_{i=1}^{n} 2^{i-1} \left[ b_{i} - \frac{k_{b}}{N-1} \right]^{2}} \bullet \sigma_{\frac{I_{RGK}}{I_{LSBX}}}$$



It can be shown that the maximum  $INL_b$  occurs at b = <011....11111> or b = <100....0000>



Substituting b=<1000....000>

$$\sigma_{INL_{b=<1000..0>}} = \sqrt{2^{n-1} \left[ 1 - \frac{N/2}{N-1} \right]^2 + \sum_{i=1}^{n-1} 2^{i-1} \left[ \frac{N/2}{N-1} \right]^2} \bullet \sigma_{\frac{I_{RGK}}{I_{LSBX}}}$$

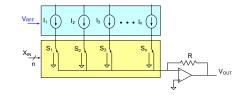
This simplifies to

$$\sigma_{INL_{b=<1000..0>}} = \sqrt{2^{n-1} \left[ 1 - \frac{N/2}{N-1} \right]^2} + \sum_{i=1}^{n-1} 2^{i-1} \left[ \frac{N/2}{N-1} \right]^2 \bullet \sigma_{\frac{I_{RGK}}{I_{LSBX}}}$$

This can be expressed as

$$\sigma_{INL_{b}} = <1000..0> = \sqrt{\frac{N}{2} \left[1 - \frac{N/2}{N-1}\right]^{2}} + \left(\frac{N}{2} - 1\right) \left[\frac{N/2}{N-1}\right]^{2} \bullet \sigma_{\frac{I_{RGK}}{I_{LSBX}}}$$

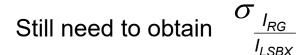
$$\sigma_{INL_{b=<1000.0>}} = \sqrt{\frac{N}{2} \left[1 - \frac{N/2}{N-1}\right]^2} + \left(\frac{N}{2} - 1\right) \left[\frac{N/2}{N-1}\right]^2} \bullet \sigma_{\frac{I_{RGK}}{I_{LSBA}}}$$



$$\sigma_{\text{INL}_{\text{MAX}}} \cong \sigma_{\text{INL}_{\text{b=<1,0,...0>}}} \cong \frac{\sqrt{N}}{2} \sigma_{\frac{I_{\text{RG}}}{I_{\text{LSBX}}}}$$

#### Note this is the same result as obtained for the unary DAC

But closed form expressions do not exist for the INL of this DAC since the INL is an order statistic



Unary Weighted and Binary Weighted

$$\sigma_{INL_{MAX}} \cong \sigma_{INL_{b=<1,0,\dots0>}} \cong \frac{\sqrt{N}}{2} \sigma_{I_{RG}}$$

Note this is the same result as obtained for the unary DAC

Since INL is the max  $|INK_k|$  is  $INL_{MAX}$  the same as INL? No!

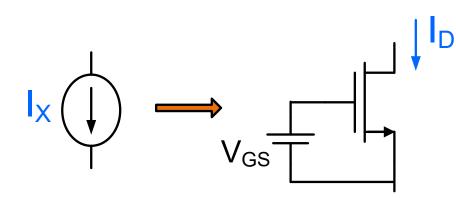
Since  $\sigma_{INL_{MAX}}$  is about the same for the Unary weighted structure and the Binary weighted structure, is the performance of both about the same? No, DNL is much different !

Instead of bundling unary current sources, could we simply take multiple outputs on a current mirror to generate the binary weighted currents ? This could be done in such a way that the area increases linearly rather than geometrically with the number of bits – much like with an R-2R DAC

Yes - but!Will $\sigma_{INL_{MAX}} \cong \frac{\sqrt{N}}{2} \sigma_{\frac{I_{RG}}{I_{LSBX}}}$  also hold if we do not bundle unary currentsources to obtain the binary current sources?

No,  $\sigma$  will be much different !

V<sub>OUT</sub>



Simple Square-Law MOSFET Model Usually Adequate for static Statistical Modeling

Assumption: Layout used to marginalize gradient effects, contact resistance and drain/source resistance neglected

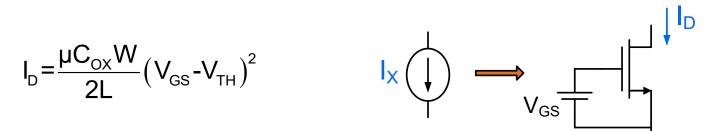
$$I_{D} = \frac{\mu C_{OX} W}{2L} (V_{GS} - V_{TH})^{2}$$

Random Variables: { $\mu$ , C<sub>OX</sub>, V<sub>TH</sub>, W, L }

Thus  $I_D$  is a random variable

From previous analysis, need:

$$\sigma_{rac{
m I_D}{
m I_{DN}}}$$



Random Variables: { $\mu$ , C<sub>OX</sub>, V<sub>TH</sub>, W, L } Thus I<sub>D</sub> is a random variable

Will <u>assume</u> { $\mu$ , C<sub>OX</sub>, V<sub>TH</sub>, W, L } are uncorrelated

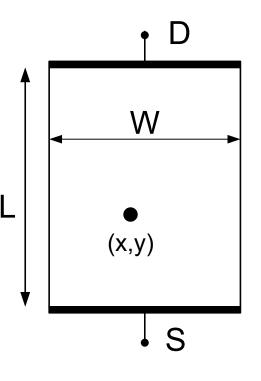
This is not true :  $T_{OX}$  is a random variable that affects both  $V_{TH}$  and  $C_{OX}$ 

- This assumption is widely used and popularized by Pelgrom
- It is also implicit in the statistical model available in simulators such as SPECTRE
- Statistical information about T<sub>ox</sub> often not available
- Drenen and McAndrew (NXP) published several papers that point out limitations
- Would be better to model physical parameters rather than model parameters but more complicated
- Statistical analysis tools at NXP probably have this right but not widely available
- Assumption simplifies analysis considerably
- Error from neglecting correlation is usually quite small but don't know how small

Model parameters are position dependent

$$I_{D} = \frac{\mu C_{OX} W}{2L} (V_{GS} - V_{TH})^{2}$$

 $\mu(x,y), \quad C_{OX}(x,y), \quad V_{TH}\left(x,y\right)$ 



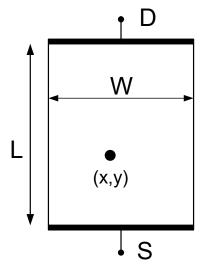
Model parameters are position dependent

$$I_{D} = \frac{\mu C_{OX} W}{2L} (V_{GS} - V_{TH})^{2}$$

Assume that model parameters can be modeled as a position-weighted integral

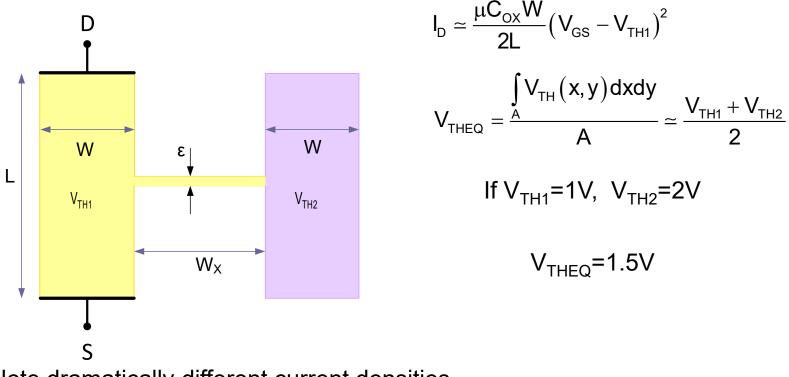
$$\mu = \frac{\int_{A}^{A} \mu(x, y) dx dy}{A}$$
$$\mu = \frac{\int_{A}^{A} C_{ox}(x, y) dx dy}{A}$$
$$C_{ox} = \frac{\int_{A}^{A} V_{TH}(x, y) dx dy}{A}$$
$$V_{TH} = \frac{\int_{A}^{A} V_{TH}(x, y) dx dy}{A}$$

Reasonably good assumption if current density is constant



Assume that model parameters can be modeled as a position-weighted integral

As seen for resistors, this model is not good if current density is not constant



Note dramatically different current densities

But reasonably good assumption if current density is constant

$$I_{\rm D} = \frac{\mu C_{\rm OX} W}{2L} (V_{\rm GS} - V_{\rm TH})^2$$

Model parameters characterized by following equations

$$\mu = \mu_N + \mu_R$$
$$V_{TH} = V_{THN} + V_{THR}$$
$$C_{OX} = C_{OXN} + C_{OXR}$$
$$L = L_N + L_R$$
$$W = W_N + W_R$$

Neglecting random part of W and L which are usually less important

$$I_{D} = \frac{(\mu_{N} + \mu_{R})(C_{OXN} + C_{OXR})W}{2L} (V_{GS} - V_{THN} - V_{THR})^{2}$$

$$I_{D} = \frac{(\mu_{N} + \mu_{R})(C_{OXN} + C_{OXR})W}{2L} (V_{GS} - V_{THN} - V_{THR})^{2}$$

This appears to be a highly nonlinear function of random variables !!

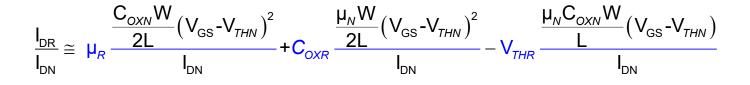
Will now linearize the relationship between  $I_D$  and the random variables Since the random variables are small, we can do a Taylor's series expansion and truncate after first-order terms to obtain

$$I_{\text{D}} \cong \frac{\mu_{\text{N}}C_{\text{OXN}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + \mu_{\text{R}} \frac{C_{\text{OXN}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXR}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}C_{\text{OXN}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXR}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}C_{\text{OXN}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXR}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}C_{\text{OXN}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXR}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}C_{\text{OXN}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXR}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}C_{\text{OXN}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXN}} \frac{\mu_{\text{N}}W}{2L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 - V_{\text{THR}} \frac{\mu_{\text{N}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXN}} \frac{\mu_{\text{N}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXN}} \frac{\mu_{\text{N}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{OXN}} \frac{\mu_{\text{N}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text{THN}} \frac{\mu_{\text{N}}W}{L} \left(V_{\text{GS}} - V_{\text{THN}}\right)^2 + C_{\text$$

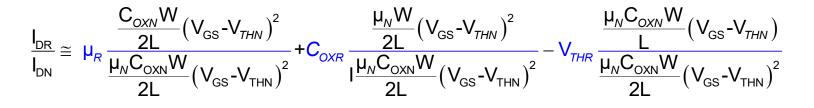
This is a linearization of  $I_{D}$  in the random variables  $\mu_{R},\,C_{OXR},\,and\,V_{THR}$ 

$$\begin{split} I_{\text{DR}} &\cong \ \mu_{R} \frac{C_{\text{OXN}} W}{2L} \left( V_{\text{GS}} - V_{\text{THN}} \right)^{2} + C_{\text{OXR}} \frac{\mu_{N} W}{2L} \left( V_{\text{GS}} - V_{\text{THN}} \right)^{2} - V_{\text{THR}} \frac{\mu_{N} C_{\text{OXN}} W}{L} \left( V_{\text{GS}} - V_{\text{THN}} \right) \\ \frac{I_{\text{DR}}}{I_{\text{DN}}} &\cong \ \mu_{R} \frac{\frac{C_{\text{OXN}} W}{2L} \left( V_{\text{GS}} - V_{\text{THN}} \right)^{2}}{I_{\text{DN}}} + C_{\text{OXR}} \frac{\frac{\mu_{N} W}{2L} \left( V_{\text{GS}} - V_{\text{THN}} \right)^{2}}{I_{\text{DN}}} - V_{\text{THR}} \frac{\frac{\mu_{N} C_{\text{OXN}} W}{L} \left( V_{\text{GS}} - V_{\text{THN}} \right)}{I_{\text{DN}}} \end{split}$$

Could easily include  $L_R$  and  $W_R$  but usually not important unless lots of perimeter



$$I_{DN} = \frac{\mu_N C_{OXN} W}{2L} (V_{GS} - V_{THN})^2$$



$$\frac{I_{DR}}{I_{DN}} \cong \frac{\mu_{R}}{\mu_{N}} + \frac{C_{OXR}}{C_{OXN}} - \frac{2V_{THR}}{(V_{GS} - V_{THN})}$$

Thus

$$\sigma_{\frac{l_{DR}}{l_{DN}}} = \sqrt{\sigma_{\frac{\mu_R}{\mu_N}}^2 + \sigma_{\frac{C_{OXR}}{C_{OXN}}}^2 + 4\left(\frac{V_{THN}}{V_{GS} - V_{THN}}\right)^2 \sigma_{\frac{V_{THR}}{V_{THN}}}^2} \qquad \text{or} \quad \sigma_{\frac{l_{DR}}{l_{DN}}} = \sqrt{\sigma_{\frac{\mu_R}{\mu_N}}^2 + \sigma_{\frac{C_{OXR}}{C_{OXN}}}^2 + \left(\frac{2}{V_{GS} - V_{THN}}\right)^2 \sigma_{\frac{V_{THR}}{V_{THR}}}^2}$$

$$\sigma_{\frac{\log}{\log}} = \sqrt{\sigma_{\frac{\log}{\log}}^{2} + \sigma_{\frac{C_{OXP}}{C_{OXN}}}^{2} + 4\left(\frac{V_{THN}}{V_{GS} - V_{THN}}\right)^{2} \sigma_{V_{THP}}^{2}} \quad \text{or} \quad \sigma_{\frac{\log}{\log}} = \sqrt{\sigma_{\frac{\log}{\log}}^{2} + \sigma_{\frac{C_{OXP}}{U_{N}}}^{2} + \left(\frac{2}{V_{GS} - V_{THN}}\right)^{2} \sigma_{V_{THP}}^{2}}$$

$$It \text{ will be assumed that} \quad \sigma_{\frac{\log}{\mu_{N}}}^{2} = \frac{A_{\mu}^{2}}{WL} \quad \text{where } A_{\mu}, A_{Cox}, A_{VT0} \text{ are Pelgrom process parameters}$$

$$\sigma_{\frac{\log}{Q}}^{2} = \frac{A_{VT0}^{2}}{WL} \quad \sigma_{\frac{\log}{Q}}^{2} = \frac{A_{\mu}^{2}}{WL}$$

$$Define \quad A_{\mu} = \sqrt{A_{\mu}^{2} + A_{Cox}^{2} + \frac{4}{V_{EB}^{2}}} A_{VT0}^{2}$$

$$Thus \quad \sigma_{\frac{\log}{Q}}^{2} = \frac{1}{\sqrt{WL}} \sqrt{A_{\mu}^{2} + A_{Cox}^{2} + \frac{4}{V_{EB}^{2}}} A_{VT0}^{2} \quad \text{Often only } A_{\beta} \text{ is available}$$

$$\sigma_{\frac{I_{DR}}{I_{DN}}} = \frac{1}{\sqrt{WL}} \sqrt{A_{\beta}^2 + \frac{4}{V_{EB}^2} A_{VT0}^2}$$

Gate area: A=WL

- Standard deviation decreases with  $\sqrt{A}$
- Large V<sub>EB</sub> reduces standard deviation
- Operating near cutoff results in large mismatch
- Often threshold voltage variations dominate mismatch

$$\sigma_{\underline{I_{DR}}\atop{\overline{I_{DN}}}}\cong \frac{2}{V_{EB}\sqrt{WL}}A_{VT0}$$



# Stay Safe and Stay Healthy !

## End of Lecture 10